

3 - Modèle linéaire

Résumé

La fiche contient le matériel nécessaire pour une séance de travaux dirigés sur S-PLUS consacrée au modèle linéaire. Elle illustre le cours de J.D. Lebreton, en particulier la régression simple, l'analyse de variance et de covariance et introduit au modèle linéaire multiplicatif.

Plan

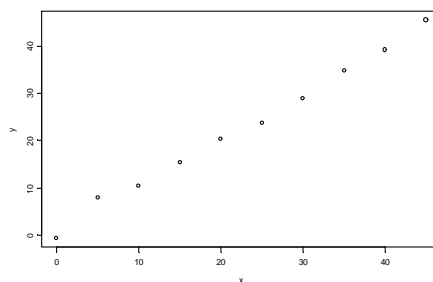
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1 - Régression simple

Lancer S-PLUS et s'assurer qu'on est bien dans le dossier de travail désiré :
Working data will be in D:\Data\DEA1_Data

Implanter le premier exemple proposé par Tomassone R., Charles-Bajard S. & Bellanger L. (1998) Introduction à la planification expérimentale, DEA « Analyse et modélisation des systèmes biologiques »:

```
> y
[1] -0.6  7.9 10.5 15.4 20.3 23.8 28.8 34.7 39.1 45.4
> x<-seq(from=0,to=45,by=5)
> x
[1]  0  5 10 15 20 25 30 35 40 45
> plot(x,y)
```



lm

```
> ?lm
```

DESCRIPTION

Returns an object of class "lm" or "mlm" that represents a fit of a linear model.

USAGE

```
lm(formula, data=<<see below>>, weights=<<see below>>,
    subset=<<see below>>, na.action=na.fail, method="qr", model=F,
    x=F, y=F, contrasts=NULL, ...)
```

REQUIRED ARGUMENTS

formula a formula object, with the response on the left of a ~ operator, and the terms, separated by + operators, on the right.

```
> lm(y~x)
Call:
lm(formula = y ~ x)
```

```
Coefficients:
(Intercept)      x
  0.7909  0.9662
```

```
Degrees of freedom: 10 total; 8 residual
```

```
Residual standard error: 1.164
```

Un modèle linéaire est un objet

```
> lm1<-lm(y~x)
> lm1
Call:
lm(formula = y ~ x)

Coefficients:
(Intercept)      x
  0.7909  0.9662

Degrees of freedom: 10 total; 8 residual
Residual standard error: 1.164
```

lm1 est de la classe lm

```
> class(lm1)
[1] "lm"
```

La classe lm est une sous-classe de la classe list

```
> is.list(lm1)
[1] T
```

lm1 est une collection de 11 composantes

```
> length(lm1)
[1] 11
> names(lm1)
 [1] "coefficients"  "residuals"      "fitted.values"  "effects"
 [5] "R"             "rank"           "assign"         "df.residual"
 [9] "contrasts"    "terms"         "call"
```

Noms et numéros des composantes de lm1

```
> lm1[[1]]
(Intercept)      x
  0.7909  0.9662
> lm1$coefficients
(Intercept)      x
  0.7909  0.9662
> lm1[[2]]
  1      2      3      4      5      6      7      8      9
-1.391 2.278 0.04727 0.1164 0.1855 -1.145 -0.9764 0.09273 -0.3382
 10
 1.131
> lm1$residuals
  1      2      3      4      5      6      7      8      9
-1.391 2.278 0.04727 0.1164 0.1855 -1.145 -0.9764 0.09273 -0.3382
 10
 1.131
```

Le calcul est possible sur les composantes

```
> 2*lm1[[1]]
(Intercept)      x
  1.582  1.932
```

Fonctions génériques : summary

```
> summary(y)
Min. 1st Qu. Median Mean 3rd Qu. Max.
-0.6   11.7    22  22.5   33.2  45.4

> summary(lm1)
```

```

Call: lm(formula = y ~ x)
Residuals:
    Min       1Q   Median       3Q      Max
-1.39 -0.817   0.07  0.168  2.28

Coefficients:
            Value Std. Error t value Pr(>|t|)
(Intercept)  0.791   0.684     1.156   0.281
            x   0.966   0.026    37.691   0.000

Residual standard error: 1.16 on 8 degrees of freedom
Multiple R-Squared: 0.994
F-statistic: 1420 on 1 and 8 degrees of freedom, the p-value is 2.69
e-010

Correlation of Coefficients:
(Intercept)
x -0.843

```

L'ordonnée à l'origine n'est pas significativement non nulle :

```

> lm2<-lm(y~-1+x)
> lm2
Call:
lm(formula = y ~ -1 + x)

Coefficients:
            x
    0.9912

Degrees of freedom: 10 total; 9 residual
Residual standard error: 1.186
> summary(lm2)

Call: lm(formula = y ~ -1 + x)
Residuals:
    Min       1Q   Median       3Q      Max
-0.979 -0.587   0.243   0.574   2.94

Coefficients:
            Value Std. Error t value Pr(>|t|)
x  0.991   0.014     70.560   0.000

Residual standard error: 1.19 on 9 degrees of freedom
Multiple R-Squared: 0.998
F-statistic: 4980 on 1 and 9 degrees of freedom, the p-value is 1.17
e-013

```

Fonctions génériques : plot

```
> ?plot
```

DESCRIPTION

Creates a plot on the current graphics device.

This function is generic (see Methods); method functions can be written to handle specific classes of data. Classes which already have methods for this function include:

data.frame, design, factor, formula, gam, glm, lm, loess, preplot.gam, preplot.loess, profile, stl, surv.fit, times, tree, tree.sequence.

USAGE

```
plot(x, ...)
```

REQUIRED ARGUMENTS

x an S-PLUS object.

```
> ?plot.lm
```

DESCRIPTION

Creates a set of plots suitable for assessing a fitted linear model of class "lm".

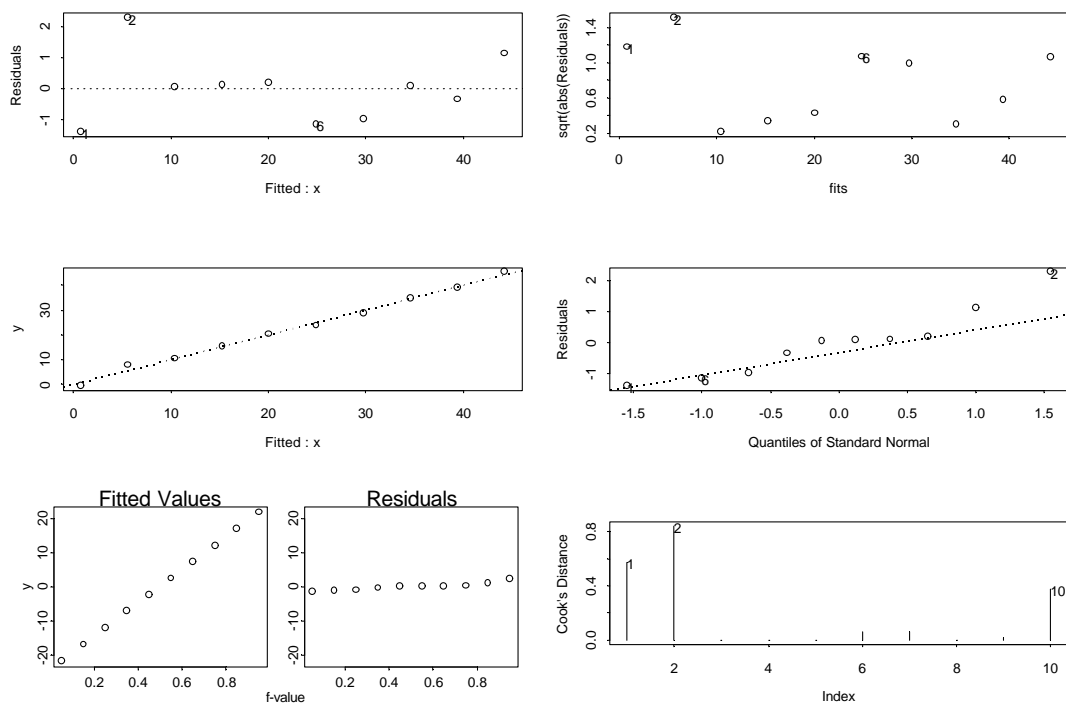
USAGE

```
plot.lm(lm.obj, residuals = NULL, smooths = F, rugplot = F,  
id.n = 3, ask = F, ...)
```

REQUIRED ARGUMENTS

lm.obj an lm object.

```
> plot(lm1) Que se passe t'il ?  
> plot(lm1,ask=T)  
> par(mfrow=c(3,2)) Pourquoi ?  
> plot(lm1)
```



Graphique standard associé à un modèle linéaire

- 1) Résidus en fonction des valeurs prédites
- 2) Racine des valeurs absolues des résidus en fonction des valeurs prédites
- 3) Valeurs observées en fonction des valeurs prédites
- 4) Graphique quantile-quantile normal des résidus (normalité des résidus). N.B. Chacun des graphiques proposés est issu d'une recherche approfondie. Le qq-plot est de Wilk M.B. & Gnanadesikan R. (1968). *Probability plotting methods for the analysis of data*. Biometrika, 55, 1-17 validé par Cleveland W.S. (1994) *The elements of graphing data*. Hobart Press, Summit, New Jersey, p. 143. Les modes de lecture sont décrits dans des ouvrages célèbres comme Tuckey J.W. (1977) *Exploratory data analysis*, Adison-Wesley, Reading, Massachussets. Ici, les résidus sont sur-dispersés

par rapport à une loi normale (cf. du Toit S.H.C., Steyn A.G.W. & Stumpf R.H. (1986) Graphical Exploratory data analysis, Springer-Verlag, , New-York, p. 49). Ouvrages classiques : Chambers J.M., Cleveland W.S., Kleiner B. & Tukey P.A. (1983) Graphical methods for data analysis, Wadsworth, Belmont, California. Cleveland W.S. (1993) Visualizing data, Hobart Press, Summit, New Jersey.

5) graphique r-f (r pour residuals, f pour fitted). A gauche, en abscisse le rang des valeurs prédites sur [0,1], en ordonnée les valeurs prédites centrées (fonction de répartition inversée des prédictions). A droite à la même échelle en abscisse le rang des résidus sur [0,1], en ordonnée les valeurs observées des résidus (fonction de répartition inversée des résidus). Le couple permet de comparer l'étendue de la distribution des observations qu'on espère beaucoup plus grande que celle des résidus. Ce graphe exprime le rapport variance expliquée - variance résiduelle.

6) Graphe des distances de Cook. Donne pour chacun des points de mesure la distance entre les paramètres estimés par la régression avec et sans ce point. Si l'importance du rôle de chaque point est concentré sur quelques valeurs, la régression n'est pas bonne (prise en compte de points aberrants). Voir Cook, R. D. and Weisberg, S. (1982). Residuals and Influence in Regression. Chapman and Hall, New York.

```
> cor(x,y)
[1] 0.997
> cor(x,y)*cor(x,y)
[1] 0.994
```

On peut refaire l'expérience :

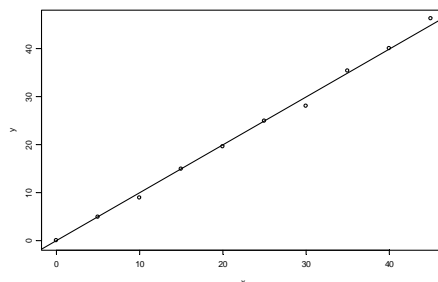
```
> x
[1] 0 5 10 15 20 25 30 35 40 45
```

norm

```
> e<-rnorm(10)
> e
[1] 0.008629 -0.038239 -1.016802 -0.132446 -0.360349 -0.033747
[7] -1.883161 0.336839 -0.000354 1.206677
```

Calcul vectoriel

```
> y<-x+e
> y
[1] 0.008629 4.961761 8.983198 14.867554 19.639651 24.966253
[7] 28.116839 35.336839 39.999646 46.206677
> par(mfrow=c(1,1)) (Sinon que se passe t'il ?)
> plot(x,y)
> abline(0,1)
```



abline

```
> ?abline
> abline(lm(y~x)) est-ce possible ?
> abline(lm(y~-1+x)) est-ce possible ?
```

2 -Analyse de variance

Reprendre l'exemple introduit dans la fiche 1 (p. 3).

lm, anova

La richesse dépend de l'heure :

```
> lm1<-lm(ric~heu.fac)
> anova(lm1)
Analysis of Variance Table

Response: ric

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
heu.fac  1      3071    3071   208.7    0
Residuals 1313      19325      15
```

La richesse dépend de la semaine :

```
> lm2<-lm(ric~heu.fac+sem.fac)
> anova(lm2)
Analysis of Variance Table

Response: ric

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
heu.fac  1      3071    3071   293.8    0
sem.fac  51      6133     120    11.5    0
Residuals 1262      13192     10
```

La richesse dépend de la station :

```
> lm3<-lm(ric~heu.fac+sem.fac+sta.fac)
> anova(lm3)
Analysis of Variance Table

Response: ric

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
heu.fac  1      3071    3071   396.6    0
sem.fac  51      6133     120    15.5    0
sta.fac  13      3519     271    35.0    0
Residuals 1249      9673      8
```

Régression polynomiale :

```
> lm4<-lm(ric~heu.fac+poly(sem,2)+sta.fac)
> anova(lm4)
Analysis of Variance Table

Response: ric

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
heu.fac  1      3071    3071   338.5    0
poly(sem, 2)  2      4025    2012   221.8    0
sta.fac  13      3523     271    29.9    0
Residuals 1298      11777      9
> lm5<-lm(ric~heu.fac+poly(sem,3)+sta.fac)
```

```
> anova(lm5)
Analysis of Variance Table
```

Response: ric

Terms added sequentially (first to last)

	Df	Sum of Sq	Mean Sq	F Value	Pr(F)
heu.fac	1	3071	3071	356.6	0
poly(sem, 3)	3	4614	1538	178.6	0
sta.fac	13	3541	272	31.6	0
Residuals	1297	11170	9		

```
> anova(lm4,lm5)
Analysis of Variance Table
```

Response: ric

	Terms	Resid. Df	RSS	Test Df
1	heu.fac + poly(sem, 2) + sta.fac	1298	11777	
2	heu.fac + poly(sem, 3) + sta.fac	1297	11170	1 vs. 2 1

	Sum of Sq	F Value	Pr(F)
1			
2	607.2	70.5	1.11e-016

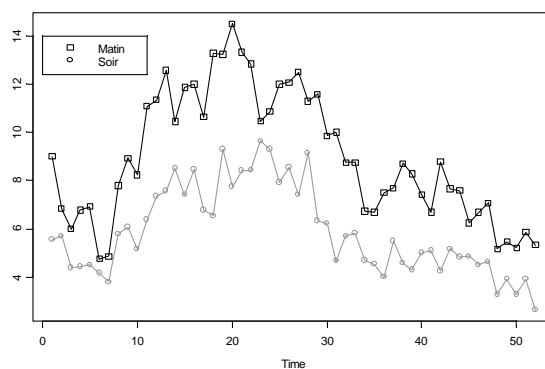
Interactions :

```
> lm6<-lm(ric~heu.fac*poly(sem,3)*sta.fac)
> anova(lm6)
Analysis of Variance Table
```

Response: ric

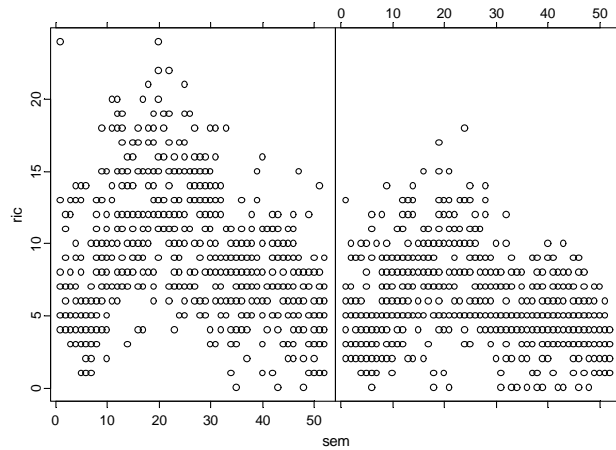
Terms added sequentially (first to last)

	Df	Sum of Sq	Mean Sq	F Value	Pr(F)
heu.fac	1	3071	3071	431.3	0.0000
poly(sem, 3)	3	4614	1538	216.0	0.0000
sta.fac	13	3541	272	38.2	0.0000
heu.fac:poly(sem, 3)	3	200	67	9.4	0.0000
heu.fac:sta.fac	13	509	39	5.5	0.0000
poly(sem, 3):sta.fac	39	1671	43	6.0	0.0000
heu.fac:poly(sem, 3):sta.fac	39	222	6	0.8	0.8047
Residuals	1203	8567	7		

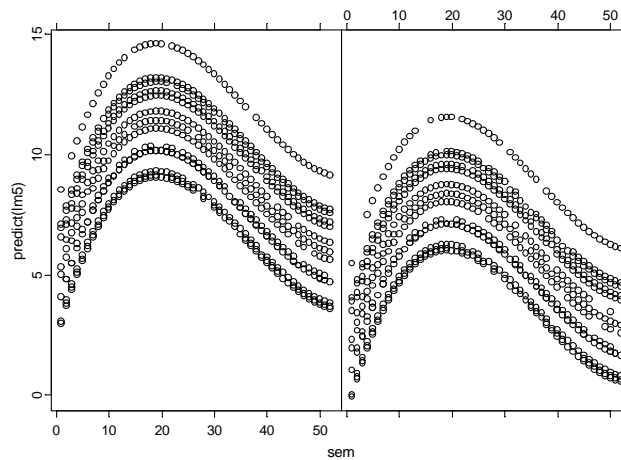


On peut mettre la même courbe pour le matin et le soir (à une constante près).

```
> coplot(ric~sem|heu.fac,show.given = F)
```

```
> coplot(predict(lm5)~sem|heu.fac, show.given = F)
```



Modèle additif sans interaction

Les modèles sont des objets

```
> lm5
Call:
lm(formula = ric ~ heu.fac + poly(sem, 3) + sta.fac)

Coefficients:
(Intercept) heu.fac poly(sem, 3)1 poly(sem, 3)2 poly(sem, 3)3
 7.407 -1.521 -29.19 -56.6 24.66
sta.fac1 sta.fac2 sta.fac3 sta.fac4 sta.fac5 sta.fac6 sta.fac7
-1.572 0.3091 -0.2556 0.7351 -0.2176 0.2248 -0.327
sta.fac8 sta.fac9 sta.fac10 sta.fac11 sta.fac12 sta.fac13
-0.1317 -0.01319 -0.1864 0.03339 0.1642 0.1017

Degrees of freedom: 1315 total; 1297 residual
Residual standard error: 2.935
> summary(lm5)

Call: lm(formula = ric ~ heu.fac + poly(sem, 3) + sta.fac)
Residuals:
    Min     1Q   Median     3Q    Max
-10.1 -1.87 -0.0598  1.82 15.5

Coefficients:
              Value Std. Error t value Pr(>|t|)
(Intercept)   7.407    0.081   91.293  0.000
heu.fac      -1.521    0.081  -18.775  0.000
poly(sem, 3)1 -29.194    2.940  -9.930  0.000
poly(sem, 3)2 -56.602    2.942  -19.238  0.000
```

```

poly(sem, 3)3 24.658 2.937 8.397 0.000
sta.fac1 -1.572 0.209 -7.516 0.000
sta.fac2 0.309 0.124 2.501 0.013
sta.fac3 -0.256 0.088 -2.908 0.004
sta.fac4 0.735 0.068 10.775 0.000
...

```

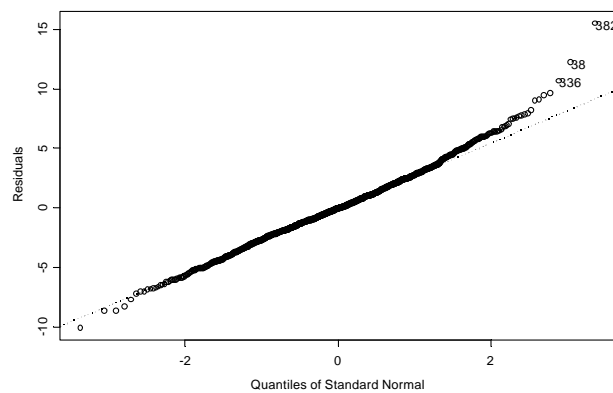
```
> plot(lm5,ask=T)
```

Make a plot selection (or 0 to exit):

```

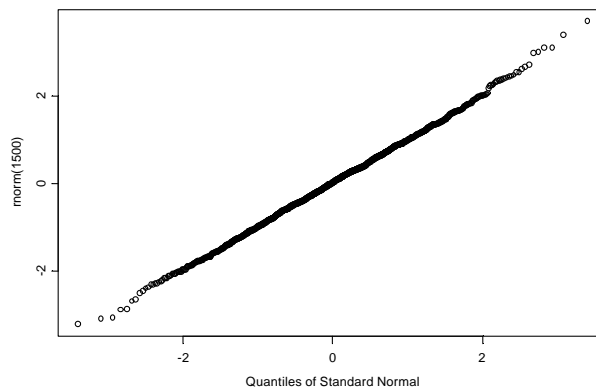
1: plot: All
2: plot: Residuals vs Fitted Values
3: plot: Sqrt of abs(Residuals) vs Fitted Values
4: plot: Response vs Fitted Values
5: plot: Normal QQplot of Residuals
6: plot: r-f spread plot
7: plot: Cook's Distances
Selection: 5

```



Graphes quantiles-quantiles : Normalité des résidus

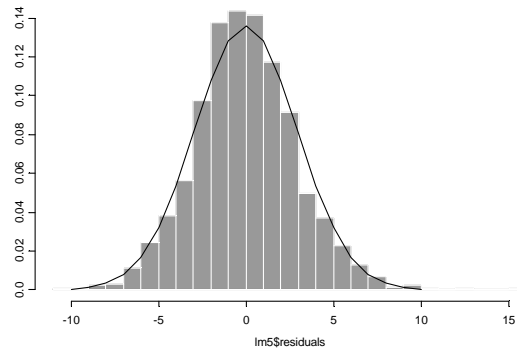
```
> qqnorm(rnorm(1500))
```



```

> ecrin[382,]
  STA SEM HEU RIC
382  5   1   1  24
> predict(lm5)[382]
 382
 8.522
> ecrin$RIC[382]-predict(lm5)[382]
 382
15.48 un point aberrant ?
> hist(lm5$residuals,proba=T,nclass=30)
> lines(seq(-10,10,1),dnorm(seq(-10,10,1),mean=0,sd=2.935))

```



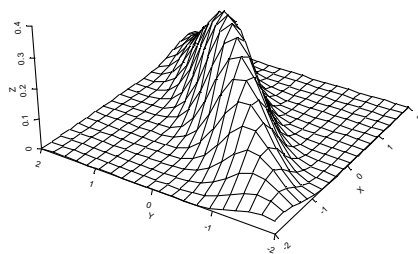
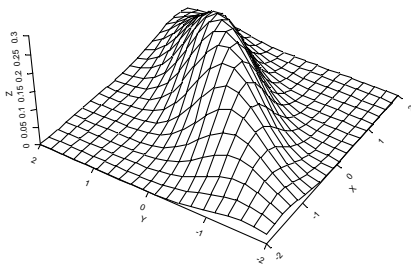
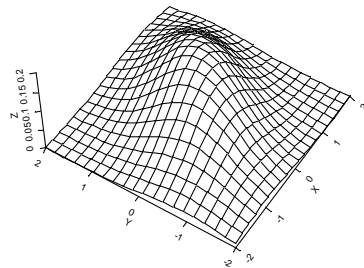
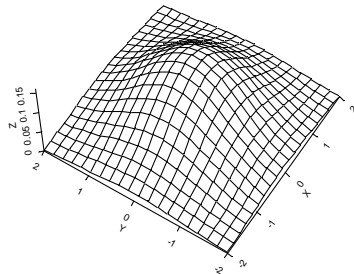
L'approche graphique des modèles statistiques : à consommer sans modération ...

3 - Densités de probabilité

3.1 - Loi de Gauss bivariée

Pour une corrélation de 0, 0.5, 0.8 et 0.9 :

```
x0<-seq(-2,2,le=20)
y0<-seq(-2,2,le=20)
xy0<-expand.grid(x=x0, y=y0)
z0<-dmvnorm(xy0,rho=0.50)
z0<-matrix(z0,nrow=20,ncol=20,byrow=T)
persp(x0,y0,z0,eye=c(-30,-20,2),box = F)
```



3.2 - Types de lignes et de caractères :

```
par(mai=c(0,0,0,0) ,cex=2)
```

```

plot(0,0,xlim=c(0,20),ylim=c(0,6),type="n",axes=F)
for (i in 1:5) {
  for(j in 1:4) {
    x0<-(j-1)*5+1
    y0<-6-i
    i0<-i+(j-1)*5-1
    text(x0,y0,paste(i0))
    points(x=x0+1,y=y0,pch=i0)
    par(lty=i0)
    segments(x0+2,y0,x0+4,y0)
  }
}

```

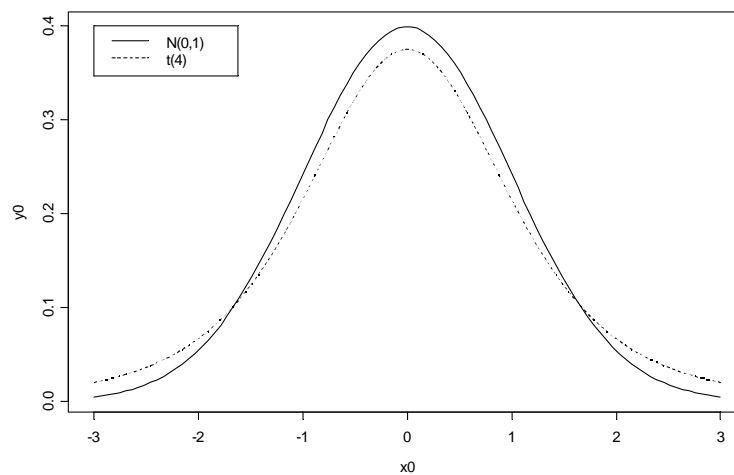
0	□	——	5		-----	10	⊕	15	■	-----
1	○	——	6	▽	-----	11	⊗	-----	16	●
2	△	7	⊠	-----	12	⊞	-----	17	▲	——
3	+	-----	8	*	13	⊠	-----	18	◆
4	×	-----	9	+	——	14	⊠	-----	19	▼	-----

3.3 - Loi de Student

```

x0<-seq(-3,3,le=100)
y0<-dnorm(x0)
z0<-dt(x0,df=4)
plot(x0,y0,type="n")
lines(x0,y0,lty=1)
lines(x0,z0,lty=8)
legend(-3, 0.4, c("N(0,1)", "t(4)"), lty = c(1,8))

```



```

> qt(df = 4, c(0.025, 0.975))
[1] -2.776 2.776
> qnorm(c(0.025, 0.975))
[1] -1.96 1.96

```

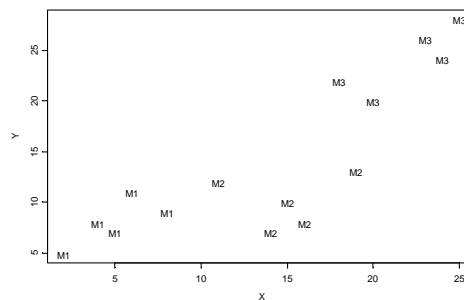
4 - Analyse de covariance

```
> M<-rep(c("M1","M2","M3"),c(5,5,5))
> M
 [1] "M1" "M1" "M1" "M1" "M1" "M2" "M2" "M2" "M2" "M2" "M3" "M3"
[13] "M3" "M3" "M3"
> X
 [1]  2  4  5  8  6 14 16 15 19 11 20 18 23 25 24
> Y
 [1]  5  8  7  9 11  7  8 10 13 12 20 22 26 28 24
```

X niveau de départ, Y niveau d'arrivée, M méthode d'enseignement.

```
> covjdl<-cbind.data.frame(X,Y,M)
> names(covjdl)<-c("X","Y","M")
> covjdl
   X  Y  M
 1  2  5 M1
 2  4  8 M1
 3  5  7 M1
...
13 23 26 M3
14 25 28 M3
15 24 24 M3

> plot(X,Y,type="n")
> text(X,Y,M)
```

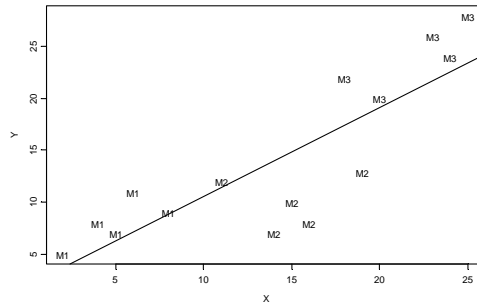


4.1 - Une seule droite de régression

```
> lm1<-lm(Y~X)
> anova(lm1)
Analysis of Variance Table

Response: Y

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value    Pr(F)
     X  1      599      599  31.53 0.00008407
Residuals 13      247       19
> abline(lm1)
```

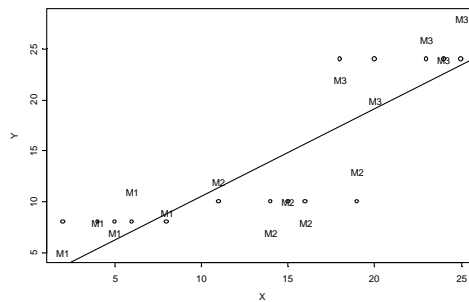


4.2 - L'effet du facteur

```
> lm2<-lm(Y~M)
> anova(lm2)
Analysis of Variance Table
```

Response: Y

```
Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value    Pr(F)
     M  2      760    380.0   53.02 1.103e-006
Residuals 12      86     7.2
> points(X,predict(lm2))
```



4.3 - Droites parallèles :

```
> lm3<-lm(Y~M+X)
> anova(lm3)
Analysis of Variance Table
```

Response: Y

```
Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value    Pr(F)
     M  2      760.0   380.0   72.58 0.00000
     X  1       28.4    28.4    5.43 0.03991
Residuals 11      57.6     5.2
```

```
> lm4<-lm(Y~X+M)
> anova(lm4)
Analysis of Variance Table
```

Response: Y

```
Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value    Pr(F)
     X  1      599.0   599.0  114.4 0.0000004
     M  2      189.4    94.7   18.1 0.0003329 Droites non égales
```

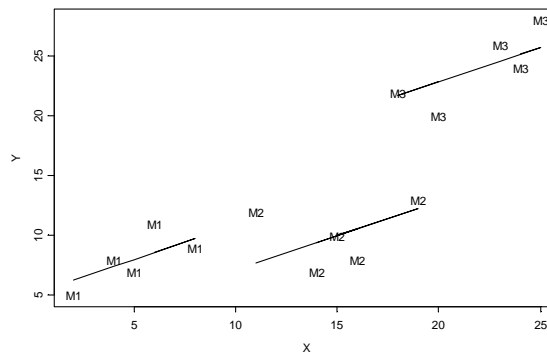
```

Residuals 11      57.6      5.2

> predict(lm3)
   1    2 3    4    5    6    7 8    9    10   11   12
6.295 7.432 8 9.705 8.568 9.432 10.57 10 12.27 7.727 22.86 21.73
   13   14   15
24.57 25.7 25.14
> predict(lm4)
   1    2 3    4    5    6    7 8    9    10   11   12
6.295 7.432 8 9.705 8.568 9.432 10.57 10 12.27 7.727 22.86 21.73
   13   14   15
24.57 25.7 25.14

> lines(X[M=="M1"],predict(lm3)[M=="M1"])
> lines(X[M=="M2"],predict(lm3)[M=="M2"])
> lines(X[M=="M3"],predict(lm3)[M=="M3"])

```



```

> coefficients(lm(Y[M=="M1"]~X[M=="M1"]))
(Intercept) X[M == "M1"]
   4.25      0.75
> coefficients(lm(Y[M=="M2"]~X[M=="M2"]))
(Intercept) X[M == "M2"]
   7.794    0.1471
> coefficients(lm(Y[M=="M3"]~X[M=="M3"]))
(Intercept) X[M == "M3"]
   4.588    0.8824

```

4.4 - Interaction

```

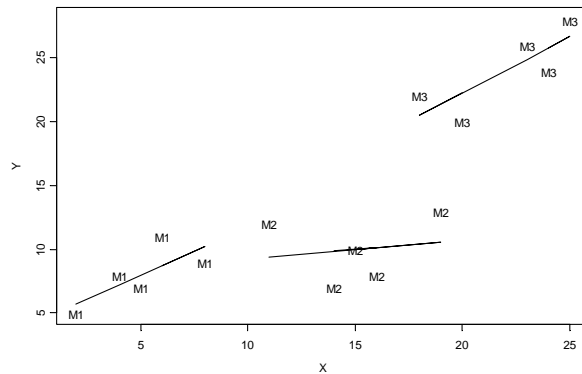
> lm5<-lm(Y~M*X)
> anova(lm5)
Analysis of Variance Table

Response: Y

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
      M  2    760.0   380.0   71.93 0.0000
      X  1     28.4    28.4    5.38 0.0456
  M:X  2     10.0     5.0    0.95 0.4220 NON SIGNIFICATIF
Residuals  9     47.5     5.3

> lines(X[M=="M1"],predict(lm5)[M=="M1"])
> lines(X[M=="M2"],predict(lm5)[M=="M2"])
> lines(X[M=="M3"],predict(lm5)[M=="M3"])

```



5 - Interaction sans répétition

16 stations météo dans les Rocheuses (4 altitudes x 4 expositions)

```
> roc
      Sud Sommet Nord Vallee
2460  88     83  74    39
2550  61     55  50    33
3000  22      9  17     0
3690 -39    -33 -39   -22

> roc.vec <- roc[, 1]
> for(i in 2:4) roc.vec <- append(roc.vec, roc[, i])
> roc.vec
[1] 88 61 22 -39 83 55 9 -33 74 50 17 -39 39 33 0 -22

> sta.vec <- names(roc)[rep(1:4, rep(4, le = 4))]
> sta.vec
[1] "Sud" "Sud" "Sud" "Sud" "Sommet" "Sommet" "Sommet"
[8] "Sommet" "Nord" "Nord" "Nord" "Nord" "Vallee" "Vallee"
[15] "Vallee" "Vallee"
> alt.vec <- row.names(roc)[rep(1:4, 4)]
> alt.vec
[1] "2460" "2550" "3000" "3690" "2460" "2550" "3000" "3690" "2460"
[10] "2550" "3000" "3690" "2460" "2550" "3000" "3690"
```

5.1 - Modèle $y_{ij} = m + a_i + b_j + e_{ij}$

```
> lm0 <- lm(roc.vec ~ sta.vec + alt.vec)
```

Propriété des plans orthogonaux

```
> anova(lm0)
Analysis of Variance Table

Response: roc.vec

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
sta.vec 3      931      310   1.94 0.194
alt.vec 3     25162     8387  52.39 0.000
Residuals 9      1441      160

> lm1 <- lm(roc.vec ~ alt.vec + sta.vec)
```



```

> anova(lm1)
Analysis of Variance Table

Response: roc.vec

Terms added sequentially (first to last)
      Df Sum of Sq Mean Sq F Value Pr(F)
alt.vec 3    25162    8387   52.39 0.000
sta.vec 3     931     310    1.94 0.194
Residuals 9    1441     160

```

```

> matrix(predict(lm0),4,4)
      [,1] [,2] [,3] [,4]
[1,] 79.12 74.62 71.63 58.625
[2,] 57.87 53.38 50.38 37.375
[3,] 20.12 15.62 12.63 -0.375
[4,] -25.12 -29.62 -32.63 -45.625

```

```

> roc-matrix(predict(lm0),4,4)
      Sud Sommet Nord Vallee
2460 8.875 8.375 2.375 -19.625
2550 3.125 1.625 -0.375 -4.375
3000 1.875 -6.625 4.375 0.375
3690 -13.875 -3.375 -6.375 23.625

```

5.2 - Modèle $y_{ij} = \mathbf{m}_i b_j + \mathbf{e}_{ij}$ ¹

```

> svd(roc)
$d:
[1] 192.914 12.069 7.762 3.524

$v:
      [,1] [,2] [,3] [,4]
[1,] -0.6008 0.4635 0.1312 -0.6380
[2,] -0.5444 -0.4482 0.6340 0.3174
[3,] -0.5119 0.3183 -0.5182 0.6067
[4,] -0.2841 -0.6949 -0.5588 -0.3523

$u:
      [,1] [,2] [,3] [,4]
[1,] -0.7621 0.003526 0.51932 0.3867
[2,] -0.5264 -0.281129 -0.19003 -0.7796
[3,] -0.1390 0.959054 -0.02792 -0.2452
[4,] 0.3505 -0.034198 0.83271 -0.4273

> svd1<-svd(roc)
> svd1$u[,1]
[1] -0.7621 -0.5264 -0.1390 0.3505
> u1<-svd1$u[,1]

> t(svd1$v[,1])
      [,1] [,2] [,3] [,4]
[1,] -0.6008 -0.5444 -0.5119 -0.2841
> v1<-t(svd1$v[,1])
> u1%*%v1
      [,1] [,2] [,3] [,4]
[1,] 0.45783 0.41485 0.39006 0.21648
[2,] 0.31627 0.28659 0.26946 0.14955
[3,] 0.08352 0.07568 0.07116 0.03949
[4,] -0.21055 -0.19078 -0.17938 -0.09956
> svd1$d[1]*u1%*%v1
      [,1] [,2] [,3] [,4]

```

¹ Mandel, J. (1961) Non additivity in two-way analysis of variance. *Journal of the American Statistical Association* : 65,878-888.

```
[1,] 88.32 80.03 75.25 41.762
[2,] 61.01 55.29 51.98 28.850
[3,] 16.11 14.60 13.73 7.618
[4,] -40.62 -36.80 -34.61 -19.206
```

```
> roc-svd1$d[1]*u1%*%v1
      Sud Sommet Nord Vallee
2460 -0.3210  2.9693 -1.248 -2.762
2550 -0.0135 -0.2865 -1.983  4.150
3000  5.8880 -5.5996  3.273 -7.618
3690  1.6175  3.8049 -4.394 -2.794
```

RAPPEL

```
> roc-matrix(predict(lm0),4,4)
      Sud Sommet Nord Vallee
2460  8.875  8.375  2.375 -19.625
2550  3.125  1.625 -0.375 -4.375
3000  1.875 -6.625  4.375  0.375
3690 -13.875 -3.375 -6.375 23.625
```

Quel est le meilleur modèle ?

```
> sum((predict(lm0)-roc.vec)^2)/9
[1] 160.1
> sum((as.vector(svd1$d[1]*u1%*%v1)-roc.vec)^2)/9
[1] 24.26

> qqnorm(predict(lm0)-roc.vec)
> qqline(predict(lm0)-roc.vec)
> qqnorm(as.vector(svd1$d[1]*u1%*%v1)-roc.vec)
> qqline(as.vector(svd1$d[1]*u1%*%v1)-roc.vec)
```

